

MID-TERM TEST

MATH 3280 3.00

OCTOBER 30, 2020

Given name and surname: _____

Student No: _____

Signature: _____

INSTRUCTIONS:

1. Please write everything in **ink**.
2. This exam is a ‘closed book’ test, duration **60** minutes.
3. Only non-programmable calculators are permitted.
4. There are five questions (each 20%) and a bonus question (10%).

USEFUL FORMULAS:

For $x \geq 0$, $t \in [0, 1)$ and $k = 0, 1, 2, \dots$, if the uniform distribution of deaths assumption holds for the life-status (x), then the following is true

$$t+k p_x \approx (1-t)_k p_x + t_{k+1} p_x.$$

GOOD LUCK!

1. In a special mortality table with a select period of one year, the following relationships are true for all ages, $x \in \{0, 1, \dots\}$

$$\begin{aligned} 0.5q_{[x]} &= 0.25q_x \\ 0.5q_{[x]+0.5} &= 0.45q_x \end{aligned}$$

Express $p_{[x]}$ in terms of p_x . Find an expression for the full expectancy of life of the life-status $[x]$.

Sol

$$\begin{aligned} \overline{p}_{[x]} &= 0.5 p(x) \cdot 0.5 p([x]+0.5) = (1 - 0.5 q_{[x]}) (1 - 0.5 q_{[x]+0.5}) \\ &= (1 - 0.25 q_x) (1 - 0.45 q_x) \\ &= (1 - 0.25(1 - p_x)) (1 - 0.45(1 - p_x)) \\ &= (0.75 + 0.25 p_x) (0.55 + 0.45 p_x) \\ &= 0.4125 + 0.475 p_x + 0.1125 p_x^2. \end{aligned}$$

Also:

$$\overline{e}_{[60]} = \sum_{k=1}^{60} k \overline{p}_{[x+k]} = \sum_{k=1}^{60} \frac{\overline{l}_{[x+k]}}{\overline{l}_{[x]}} = \frac{\overline{l}_x}{\overline{l}_{[60]}} \sum_{k=1}^{60} \frac{\overline{l}_{x+k}}{\overline{l}_x} = \frac{\overline{l}_x}{\overline{l}_{[60]}} e_x$$

2. Using the Illustrative Life Table, compute the constant force of mortality applicable to a life aged between ages 67 and 68 exact. Then compute $0.5q_{67.25}$ using the constant force of mortality assumption.

Sol

We have

$$p_{67} = \exp \left\{ - \int_{67}^{68} p_{67} dt \right\} \Rightarrow p_{67} = \ln(p_{67}) = -\ln(1 - q_{67})$$

$$= -\ln \left(1 - \frac{1}{1600} 25.4391 \right) = 0.025768$$

Then

$$0.5q_{67.25} = 1 - 0.5p_{67.25} = 1 - 0.5p_{67}$$

$$= 1 - \exp \left\{ -0.5 \ln(p_{67}) \right\} = 0.6128$$

3. Evaluate

$$\frac{\partial}{\partial x} \overset{\circ}{e}_{x:y}$$

under the assumption of independence of the random variables $T(x)$ and $T(y)$, where $x, y > 0$.

$$\begin{aligned}
 \frac{\partial}{\partial x} \overset{\circ}{e}_{x:y} &\stackrel{(1)}{=} \frac{\partial}{\partial x} \int_0^\infty e^{px + py} dt = \int_0^\infty \frac{\partial}{\partial x} e^{px + py} dt \\
 &\stackrel{(2)}{=} \int_0^\infty \frac{\partial}{\partial x} \exp \left\{ - \int_0^t p(x+s) ds \right\} e^{py} dt \\
 &= - \int_0^\infty \exp \left\{ - \int_0^t p(x+s) ds \right\} \left(\frac{\partial}{\partial x} \int_0^t p(x+s) ds \right) e^{py} dt \\
 &\stackrel{(3)}{=} - \int_0^\infty e^{px} e^{py} \left(\int_0^t \frac{\partial}{\partial x} p(x+s) ds \right) dt \\
 &\stackrel{(4)}{=} - \int_0^\infty e^{px} e^{py} \left(p(x+t) - p(x) \right) dt \\
 &\stackrel{(5)}{=} \int_x^\infty e^{px} e^{py} (p(x) - p(x+1)) dt \\
 &\stackrel{(6)}{=} p(x) \overset{\circ}{e}_{x:y} - \int_x^\infty e^{px+py} p(x+t) dt
 \end{aligned}$$

4. Let $T'(x)$, $T'(y)$, ξ be independent random variables distributed Pareto with the respective shape parameters $\gamma'_x > 0$, $\gamma'_y > 0$, $\gamma > 0$ and the common scale parameter $\lambda > 0$. Set $T(x) = (T'(x) \wedge \xi)$ and $T(y) = (T'(y) \wedge \xi)$. (Recall that, for instance, $\mathbb{P}(\xi > t) = (1 + t/\lambda)^{-\gamma}$, $\lambda > 0$, $\gamma > 0$, $t \geq 0$.) Derive the probability $\mathbb{P}(T(x:y) \geq t)$, $t \geq 0$. Use this probability to find $\mu_{x:y}(t) = \mu((x:y) + t)$, $t > 0$.

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$$\begin{aligned} \mathbb{P}_{x:y} &= \mathbb{P}(T(x:y) \geq t) = \mathbb{P}(T'(x) \wedge T'(y) \geq t, T'(x:y) \geq t) \\ &\geq \mathbb{P}(T'(x) \geq t, T'(y) \geq t, \xi \geq t) \\ &\stackrel{(1)}{=} \mathbb{P}(T'(x) \geq t) \mathbb{P}(T'(y) \geq t) \mathbb{P}(\xi \geq t) \\ &\geq \left(1 + \frac{t}{\lambda}\right)^{-\gamma'_x} \left(1 + \frac{t}{\lambda}\right)^{-\gamma'_y} \left(1 + \frac{t}{\lambda}\right)^{-\gamma} \\ &= \left(1 + \frac{t}{\lambda}\right)^{-(\gamma + \gamma'_x + \gamma'_y)}, \quad t \geq 0. \end{aligned}$$

Also, we obtain

$$-\frac{d}{dt} \mathbb{P}_{x:y} \geq \frac{1}{\lambda} (\gamma + \gamma'_x + \gamma'_y) \left(1 + \frac{t}{\lambda}\right)^{-(\gamma + \gamma'_x + \gamma'_y + 1)}$$

and so

$$\mu_{x:y}(t) = \frac{\gamma + \gamma'_x + \gamma'_y}{\lambda} \left(1 + \frac{t}{\lambda}\right)^{-\frac{1}{2}} \frac{\gamma + \gamma'_x + \gamma'_y}{\lambda + t}, \quad t > 0.$$

5. Let the random variables $T(x)$ and $T(y)$ be independent and distributed $Uni[0, 1]$. Simplify the expression for the joint probability density function of the random variable $T(x : y)$ as much as possible.

Sol

$$\epsilon p_{xy} = \epsilon p_x \cdot \epsilon p_y = (1-\epsilon) (1-\epsilon) = 1 - 2\epsilon + \epsilon^2, \epsilon \in [0, 1]$$

Hence

$$- \frac{d}{dt} \epsilon p_{xy} = -\frac{d}{dt} (1 - 2\epsilon + \epsilon^2) = 2 - 2\epsilon = 2(1-\epsilon), \epsilon \in (0, 1)$$

which is the PDF

Also

$$\begin{aligned} & P(T(x:y) \geq t \mid T(x) < 1, T(y) < 1) \\ &= \frac{P(T(x) \geq t, T(y) \geq t, T(x) < 1, T(y) < 1)}{P(T(x) < 1, T(y) < 1)} \\ &\stackrel{4}{=} \frac{P(\epsilon \leq T(x) < 1) P(\epsilon \leq T(y) < 1)}{P(T(x) < 1, T(y) < 1)} \end{aligned}$$

$$\stackrel{2}{=} \frac{(q_x - \epsilon q_x)(q_y - \epsilon q_y)}{q_x q_y}$$

$$\stackrel{2}{=} \frac{q_x q_y - q_x \epsilon q_y - \epsilon q_x q_y + \epsilon^2 q_x q_y}{q_x q_y}$$

$$\stackrel{2}{=} \frac{q_x q_y - t q_x q_y - t q_x q_y + t^2 q_x q_y}{q_x q_y}$$

$$\stackrel{2}{=} 1 - 2t + t^2 \neq 1 - t,$$

Hence the PDF is

$$- \frac{d}{dt} (1 - 2t + t^2) = 2 - 2t, t \in [0, 1]$$

6. You are given that the function $t \mapsto \mu_x(t)$ is constant for $0 \leq t < 1$ and that $q_x = 0.16$.
 Find the value of $t > 0$ such that ${}_t p_x = 0.95$.

S6]

$${}_t p_x = \exp \left\{ - \int_0^t \mu_x(s) ds \right\} \stackrel{\text{defn}}{=} \exp \left\{ - \int_0^t \mu_x ds \right\} = e^{-\mu_x t}, \quad t \geq 0$$

$$\mu_x = 1 - q_x = 0.84 \approx e^{-\mu_x} \Leftrightarrow \mu_x = -\ln(0.84) \approx 0.17435$$

So we must solve for t

$$0.95 \approx e^{-0.17435 t}$$

$$\Leftrightarrow t \approx \frac{-1}{0.17435} \ln(0.95) \approx 0.294192$$