

# Life status

## Definition 1.1 (Random variable)

Let  $(\Omega, \mathcal{F}, \mathbf{P})$  be a probability triple where  $\Omega$  is sample space,  $\mathcal{F}$  is sigma algebra over  $\Omega$ , and  $\mathbf{P}$  is probability measure. Random variable,  $X$ , is a map  $X : \Omega \rightarrow \mathcal{A} \subseteq \mathbf{R}$ .

Remember, random variables are in fact functions, i.e., it is more appropriate to write  $X(\omega)$  with  $\omega \in \Omega$ . However, we often (from now on) omit  $(\omega)$  to make our writing simpler.

## Example 1.1 ( $\mathcal{A}$ will be sometimes finite in this course)

Let  $\Omega = \{\textit{sick}, \textit{healthy}, \textit{very healthy}\}$ , and define an r.v. as:

$$X(\omega) = \begin{cases} 20, & \omega = \textit{sick} \\ 60, & \omega = \textit{healthy} \\ 100, & \omega = \textit{very healthy} \end{cases} .$$

### Example 1.2 ( $\mathcal{A}$ does not have to be finite)

Imagine an experiment in which we throw a needle to pick up a number in the interval  $(0, 123]$ . In such a case,  $\omega$  denotes an outcome of the experiment, i.e.,  $\omega \in \Omega = \{(0, 123]\}$ . We can then define an r.v.  $Y(\omega) = \omega$ , with the domain and range being  $\mathcal{A} = \{(0, 123]\}$ .

### Note.

The example above can be easily reformulated to make  $Y$  describe a future life time of ( $x$ ) (= person age  $x$ ). Note that the maximal future life time of such a person cannot exceed 123.

### Definition 1.2 (Continuous future life time of a new born)

Let  $T(0) : \Omega \rightarrow \mathbf{R}_+ := (0, \infty)$  denote the r.v. describing the future life time of a new born child.

**Definition 1.3 (Continuous future life time of an adult)**

Let  $x > 0$ , then  $T(x) = (T(0) - x | T(0) > x) : \Omega_x \rightarrow \mathbf{R}_+$  denotes an r.v. describing the future life time of a person age  $x$ .

**Definition 1.4 (Curtate future life time of a new born)**

Let  $K(0) : \Omega \rightarrow \mathbf{Z}_+ := \{0, 1, \dots\}$  denote an r.v. describing the curtate future life time of a new born child. We have that  $K(0) := \lfloor T(0) \rfloor$ .

**Definition 1.5 (Curtate future life time of an adult)**

Let  $x \in \mathbf{Z}_+$ , then  $K(x) := (K(0) - x | K(0) \geq x) : \Omega_x \rightarrow \mathbf{Z}_+$  denotes an r.v. describing the curtate future life time of an adult age  $x$ . We again have that  $K(x) := \lfloor T(x) \rfloor$ .

Note,  $(x)$  and  $(0)$  above are not related to the concept of sample spaces. These are special cases of what we shall call 'life status'.

### Definition 1.6 (Life status)

A life status  $(u)$  is an artificially constructed life form for which the notion of life and death can be well defined.

### Example 1.3

The life status  $(u) := (0)$  dies when the new born child dies.

### Example 1.4

The status  $(u) := (x)$  dies when the person age  $x$  dies.

### Definition 1.7 ('Angle $n$ ')

For  $n > 0$ , we have that  $(u) := (\bar{n})$  is a life status that survives  $n$  time units (alive on  $[0, n)$ ) and dies thereafter (dead on  $[n, \infty)$ ).

### Example 1.5 (Simple joint life status)

The status  $(u) := (x : \bar{n})$  dies when  $(x)$  dies if the death occurs before time  $n$ , and it is always dead if otherwise.

### Definition 1.8 (Joint life status)

The life status  $(x : y)$  is called the joint life status, and it dies on the first death of either  $(x)$  or  $(y)$ .

### Note

If payments of, say, an annuity stop when  $(x : y)$  dies, then the longer living person is not get paid.

### Definition 1.9 (Last survivor life status)

The life status  $(\overline{x} : \overline{y})$  dies on the last death of either  $(x)$  or  $(y)$ .

### Note.

- What if the lives of  $(x)$  and  $(y)$  are dependent?
- The notions above can be extended to augment more than just two lives.

### Definition 1.10 (Probability of death)

The c.d.f. of a general life status ( $u$ ) is given by

$$\mathbf{P}[T(u) \leq t] = F_{T(u)}(t) \text{ or } \mathbf{P}[K(u) \leq k] = F_{K(u)}(k)$$

and it provides the probability that ( $u$ ) dies during the interval  $(0, t]$  or  $[0, k]$ .

The c.d.f.  $F$  is a map from  $R \subseteq \mathbf{R} := (-\infty, \infty)$  to  $[0, 1]$ , with  $R$  being the range of the r.v. of interest, such that

- $F$  is non-decreasing, i.e., for  $x_1, x_2$  in  $R$ , and  $x_1 \leq x_2$ , we have that  $F(x_1) \leq F(x_2)$ .
- $F$  is right continuous, i.e.,  $F(x_0+) = \lim_{x \downarrow x_0} F(x) = F(x_0)$ .
- $F(+\infty) = 1$ ,
- $F_{T(u)}(0) = 0$ , and  $F_{K(u)}(0-) = 0$ ,
- $F(x_0) = 1 - \bar{F}(x_0)$ , for  $x_0 \in R$ .

### Definition 1.11 (Probability to survive)

The d.d.f. (also the survival function) of a general life status ( $u$ ) is given by

$$\mathbf{P}[T(u) > t] = \bar{F}_{T(u)}(t) \text{ or } \mathbf{P}[K(u) > k] = \bar{F}_{K(u)}(k)$$

and it provides the probability that ( $u$ ) survives more than  $t$  time units.

The d.d.f.  $\bar{F}$  is a map from  $R \subseteq \mathbf{R} := (-\infty, \infty)$  to  $[0, 1]$ , with  $R$  being the range of the r.v. of interest, such that

- $\bar{F}$  is non-increasing, i.e., for  $X_1, x_2$  in  $R$  and  $x_1 \leq x_2$ , we have that  $\bar{F}(x_1) \geq \bar{F}(x_2)$ .
- $\bar{F}$  is right continuous, i.e.,  $\bar{F}(x_0+) = \lim_{x \downarrow x_0} \bar{F}(x) = \bar{F}(x_0)$ .
- $\bar{F}(+\infty) = 0$
- $\bar{F}_{T(u)}(0) = 1$ , and  $\bar{F}_{K(u)}(0-) = 1$ ,
- $\bar{F}(x_0) = 1 - F(x_0)$ , for  $x_0 \in R$ .

## Definition 1.12

We shall denote by

$${}_t p_u := \mathbf{P}[T(u) \geq t] = \bar{F}_T(t-) \text{ and } {}_t p := \mathbf{P}[T(0) \geq t] = \bar{F}(t-)$$

the probability that  $(u)$  (and  $(0)$ ) survives at least  $t$  time units, and by

$${}_t q_u = \mathbf{P}[T(u) < t] = F_T(t-) \text{ and } {}_t q := \mathbf{P}[T(0) < t] = F(t-)$$

the probability that  $(u)$  dies before it attains the  $t$ -th time unit.

## Note.

The events  $\{T(u) \geq t\}$  and  $\{T(u) < t\}$  are disjoint and complement each other. Thus  $1 = {}_t p_u + {}_t q_u$ , for any fixed  $t \geq 0$  and a general life status  $(u)$ . Same holds for the curtate counterpart.



## Note.

Note that  ${}_t p_u$  and  ${}_t q_u$  are not, respectively, the d.d.f and the c.d.f., however for continuous r.v.'s, we readily have that

$$F(t) - F(t-) = 0 = \mathbf{P}[T(u) = t], \quad (1)$$

and thus in such a case

$${}_t q_u := \mathbf{P}[T(u) < t] \stackrel{(1)}{=} \mathbf{P}[T(u) \leq t] = F(t),$$

and

$${}_t p_u := \mathbf{P}[T(u) \geq t] \stackrel{(1)}{=} \mathbf{P}[T(u) > t] = \bar{F}(t).$$

- How do we define a continuous r.v.?
- What other r.v.'s do exist out there?

### Definition 1.13 (Atoms)

Let  $F$  be a c.d.f., then the values of  $x \in R \subseteq \mathbf{R}$  are called atoms if

$$F(x) - F(x-) > 0.$$

We shall denote by  $\mathcal{D}_F$  the set of such atoms of  $F$ , i.e.,

$$\mathcal{D}_F := \{x : F(x) - F(x-) > 0\}.$$

### Definition 1.14 (Continuous random variables)

The r.v. is continuous if  $\mathcal{D}_F = \emptyset$ . Hence

$$\sum_{x \in R} (F(x) - F(x-)) = 0.$$

In this course continuous r.v.'s will be denoted by  $T(\cdot)$ .

- Can you think of any continuous r.v.'s?

**Definition 1.15 (Discrete random variables)**

The r.v. is discrete if its range  $R$  is countable, and hence

$$\sum_{x \in R} (F(x) - F(x-)) = 1.$$

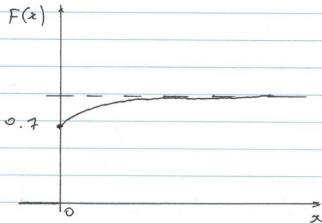
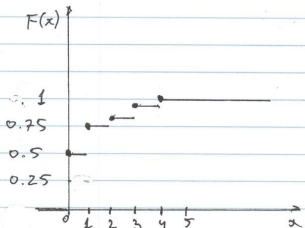
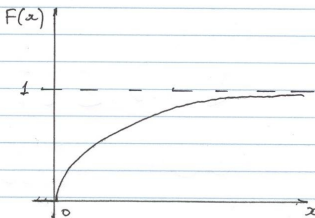
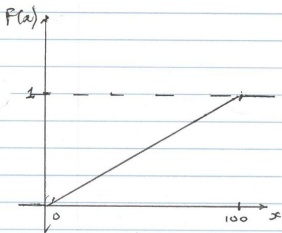
In this course discrete r.v.'s will be denoted by  $K(\cdot)$ .

**Definition 1.16 (Mixed random variables)**

The r.v. is mixed if its c.d.f. is such that

$$0 < \sum_{x \in R} (F(x) - F(x-)) < 1.$$

- We shall see one more type of random variables in this course. Any guess?



### Definition 1.17 (Absolutely continuous random variables)

The r.v. is absolutely continuous if there exists a function (density)  $f(x) > 0$ ,  $x \in \mathbf{R}$  such that, for  $a < b$ , we have that

$$F(b) - F(a) = \int_a^b f(x) dx.$$

If  $f$  exists, then it is called the probability density function (p.d.f.).

Useful absolutely continuous r.v.'s to remember:

- Exponential:  $T(0) \sim \text{Exp}(\lambda)$ ,  $\lambda > 0$ . The p.d.f. is  $f(x) = \lambda e^{-\lambda x} \mathbf{1}\{x > 0\}$ . The c.d.f. is  $F(x) = 1 - e^{-\lambda x}$ .
- Gamma:  $T(0) \sim \text{Ga}(\gamma, \alpha)$ ,  $\gamma, \alpha > 0$ . The p.d.f. is  $f(x) = e^{-\alpha x} x^{\gamma-1} \alpha^\gamma (\Gamma(\gamma))^{-1} \mathbf{1}\{x > 0\}$ . The c.d.f. is...?
- Normal:  $N(\mu, \sigma^2)$ . The p.d.f. is  $f(x) = (\sqrt{2\pi\sigma^2})^{-1} \exp\{-0.5(x - \mu)^2/\sigma^2\} \mathbf{1}\{x \in \mathbf{R}\}$

### Definition 1.18 (Force of mortality.)

For an absolutely continuous r.v.  $T(0)$  with the c.d.f.  $F$ , d.d.f.  $\bar{F}$  and p.d.f.  $f$ , and  $t > 0$ , the force of mortality is

$$\mu(t) := \frac{f_{T(0)}(t)}{\bar{F}_{T(0)}(t)}.$$

### Motivation for $\mu(\cdot)$ .

For  $T(0)$  as above, we have that

$$\begin{aligned} \mu_{T(0)}(t) &= \frac{F'(t)}{\bar{F}(t)} = \frac{1}{\bar{F}(t)} \lim_{h \downarrow 0} \frac{F(t+h) - F(t)}{h} \\ &= \frac{1}{\bar{F}(t)} \lim_{h \downarrow 0} \frac{\mathbf{P}[t < T(0) \leq t+h]}{h} \\ &= \lim_{h \downarrow 0} \frac{\mathbf{P}[t < T(0) \leq t+h \mid T(0) > t]}{h}. \end{aligned}$$

## Remember

For each age  $t > 0$ ,  $\mu(t)$  gives the value of the conditional density function of  $T(0)$  at exact age  $t$  given survival to that age.

## At home

Show that the force of mortality of  $T(u)$ , if exists, is given by

$$\mu_{T(u)}(t) = \mu(u + t).$$

## Remember

In actuarial notation, we have that

$$\mu_{T(u)}(t) = \frac{1}{t p_u} \frac{d}{dt} t q_u = -\frac{1}{t p_u} \frac{d}{dt} t p_u = \frac{1}{t+u p} \frac{d}{dt} t+u q.$$

- It is possible to find a relation between the force of mortality and the probability of survival/death.

## Proposition 1.1

Let  $T(u)$  have the force of mortality  $\mu(u + t)$ . Then

$${}_t p_u = \mathbf{P}[T(u) \geq t] = \bar{F}_{T(u)}(t) = \exp \left\{ - \int_0^t \mu(u + s) ds \right\}.$$

## Proof.

Start with the definition  $\mu(u + s) = -\bar{F}'(u + s)/\bar{F}(u + s)$ , for a fixed  $u$ , and rewrite it as

$$\mu(u + s) = -\frac{d}{ds} \ln \bar{F}(u + s).$$

Integrate both sides

$$- \int_0^t \mu(u + s) ds = \int_0^t \frac{d}{ds} \ln \bar{F}(u + s) = \ln \frac{\bar{F}(u + t)}{\bar{F}(u)},$$



Cont.

and notice that

$$\begin{aligned}
 {}_t p_u &= \mathbf{P}[T(u) \geq t] = \mathbf{P}[T(u) > t] \\
 &= \mathbf{P}[T(0) > u + t \mid T(0) > u] = \frac{\mathbf{P}[T(0) > u + t]}{\mathbf{P}[T(0) > u]} \\
 &= \frac{\bar{F}(u + t)}{\bar{F}(u)} = \frac{{}_{u+t}p}{{}_u p}.
 \end{aligned}$$

This

$$- \int_0^t \mu(u + s) ds = \ln {}_t p_u,$$

which leads to

$${}_t p_u = \exp \left\{ - \int_0^t \mu(u + s) ds \right\},$$

and hence completes the proof.



**At home.**

In the proof of Proposition 1.1, we showed that

$${}_t p_u = {}_{u+t} p / {}_u p$$

for any positive  $u$  and  $t$  and a continuous r.v.  $T(u)$ . Show that the continuity is not necessary, and the statement is still true.

**Note.**

From Proposition 1.1, we have that (change of variables),

$${}_t p_u = \exp \left\{ - \int_u^{u+t} \mu(s) ds \right\},$$

and also (put  $u = 0$ ),

$${}_t p = \exp \left\{ - \int_0^t \mu(s) ds \right\},$$

## Proposition 1.2

Let  $T(u)$  be an absolutely continuous r.v., then its p.d.f. is, for positive  $u$  and  $t$ ,

$$f_{T(u)}(t) = {}_t p_u \mu(u + t).$$

## Proof.

Use Proposition 1.1 and have that

$$-\frac{d}{dt} {}_t p_u = \exp \left\{ - \int_u^{u+t} \mu(s) ds \right\} \mu(u + t) = {}_t p_u \mu(u + t).$$

which completes the proof. □

## Note.

Put  $u = 0$ , and the p.d.f. of  $T(0)$  becomes

$$f(t) = {}_t p \mu(t).$$

### Proposition 1.3 (Property of $\mu$ .)

If  $\mu$  is a legitimate force of mortality, then we have that

$$\lim_{t \uparrow \infty} \int_U^{u+t} \mu(s) ds = \infty.$$

### Proof.

Note that  $\lim_{t \uparrow \infty} {}_t p_U = 0$  and hence  $\lim_{t \uparrow \infty} \ln {}_t p_U = -\infty$ . Then use Proposition 1.1 and have that

$$\lim_{t \uparrow \infty} \int_U^{u+t} \mu(s) ds = -\lim_{t \uparrow \infty} \ln {}_t p_U = \infty,$$

which completes the proof. □

- Proposition 1.3 means that not every function can serve as a force of mortality.

## Remember.

For a function to be a legitimate force of mortality, we must have that

- $\mu(u + t) \geq 0$  - conditional p.d.f., and
- $\lim_{t \uparrow \infty} \int_u^{u+t} \mu(s) ds = \infty$  - the accumulated force of mortality eventually becomes infinitely strong and kills.

## At home.

We said  $\mu$  has an interpretation of a conditional density, why then do we have Proposition 1.3? We know that a density integrates to one.

- Is it very practical to consider continuous r.v.'s describing future life times?

### Proposition 1.4

Let  $K(u)$  be a curtate r.v. representing the future life time of a life status  $(u)$ . The probability mass function (p.m.f.) of  $K(u)$  is given by

$$\mathbf{P}[K(u) = k] = {}_k p_u \cdot q_{u+k}.$$

### Proof.

Remember that  $K(u)$  is the highest integer in  $T(u)$ . We have that, for  $k = 0, 1, \dots$ ,

$$\begin{aligned} \mathbf{P}[K(u) = k] &= \mathbf{P}[k \leq T(u) < k + 1] = \mathbf{P}[k < T(u) \leq k + 1] \\ &= F_{T(u)}(k + 1) - F_{T(u)}(k) = {}_{k+1} q_u - {}_k q_u = {}_k p_u \cdot q_{u+k} \end{aligned}$$

Also  $\sum_{k=0}^{\infty} {}_k p_u \cdot q_{u+k} = {}_{\infty} q_u$ , and the p.m.f. is a legitimate one. This completes the proof.  $\square$

## Remember.

We shall denote the probability  $\mathbf{P}[K(u) = k]$  by  ${}_kq_u$ , that is a deferred probability of death during the time unit following the  $k$ -th one. In a similar fashion, the deferred probability to die in  $n$  time units during the consequent  $m$  time units is  ${}_{n|m}q_u$ .

- Remember that for a discrete r.v.  $K(u)$ , the probability  $\mathbf{P}[K(u) \leq k]$  is not equal to  ${}_kq_u$ . So what is the c.d.f. of  $K(u)$ ?

## Proposition 1.5

The c.d.f. of  $K(u)$  is given by, for  $k = 0, 1, \dots$ ,

$$F_{K(u)}(k) = {}_{k+1}q_u.$$

## Proof.

$$\begin{aligned}
 F_{K(u)}(k) &= \sum_{l=0}^k \mathbf{P}[K(u) = l] = \sum_{l=0}^k {}_l q_u \\
 &= \sum_{l=0}^k {}_{l+1} q_u - {}_l q_u = {}_{k+1} q_u - {}_0 q_u = {}_{k+1} q_u,
 \end{aligned}$$

which completes the proof. □

## Remember.

We assumed that the observation of survival at age  $u$  yields the same conditional distribution of survival as the hypothesis that (0) survived to age  $u$ , i.e.,

$$\mathbf{P}[T(u) > t] = \mathbf{P}[T(0) > u + t \mid T(0) > u]. \text{ Natural?}$$



## Simple life status (0).

At home.

Show that, for positive  $n$ ,  $m$ ,  $u$ ,

$${}_{n|m}q_u = {}_n p_u \cdot {}_m q_{u+n}.$$

Explain.

### Example 1.6 (Simple life status for a new born.)

Let  $(u) = (0)$ , i.e., we deal with a life status that dies when a new born child dies. In such a case,  $K(0)$  describes the future life time of the child in discrete time units (often years), and  $T(0)$  describes it continuously. For the former, we have that the p.m.f. is  ${}_k|q$  and the c.d.f. is  ${}_{k+1}q$ . Also, for the latter, we have that the p.d.f. is  ${}_t p \mu(t)$ , the c.d.f. is  ${}_t q$  and the force of mortality is  $\mu(t)$ . Here  $k = 0, 1, \dots$  and  $t > 0$ .

Simple life status  $(x)$ .Example 1.7 (Simple life status for a person age  $x$ .)

Let  $(u) = (x)$ , i.e., we deal with a life status that dies when a person age  $x$  dies. In such a case,  $K(x)$  describes the future life time of that person in discrete time units (often years), and  $T(x)$  describes it continuously. For the former, we have that the p.m.f. is  ${}_k|q_x$  and the c.d.f. is  ${}_{k+1}q_x$ . Also, for the latter, we have that the p.d.f. is  ${}_t p_x \mu(x+t)$ , the c.d.f. is  ${}_t q_x$  and the force of mortality is  $\mu(x+t)$ . Here  $k = 0, 1, \dots$ ,  $x > 0$  and  $t > 0$ .

Simple life status  $(\bar{n})$ .

## Example 1.8

Let  $(u) = (\bar{n})$ , i.e., we specialize our general life status to the one that is alive before time  $n$  and is dead thereafter. The future life time is therefore described by the r.v.  $T(\bar{n})$ , such that

$${}_tq_{\bar{n}} = \begin{cases} 0, & t < n \\ 1, & t \geq n \end{cases} \Leftrightarrow {}_tp_{\bar{n}} = \begin{cases} 1, & t < n \\ 0, & t \geq n \end{cases} .$$

$K(\bar{n}) := \lfloor T(\bar{n}) \rfloor$ , such that

$$\mathbf{P}[K(\bar{n}) = k] = \begin{cases} 1, & k = n - 1 \\ 0, & k \neq n - 1 \end{cases} ,$$

$${}_kp_{\bar{n}} = \begin{cases} 1, & k \leq n - 1 \\ 0, & k > n - 1 \end{cases} = \begin{cases} 1, & k < n \\ 0, & k \geq n \end{cases}$$

At home.

If one dollar is paid upon death of  $(\bar{n})$ , what is the a.p.v. of the payment given an interest rate  $i$ ?

# A simple joint life status $(x : \bar{n})$ .

## Example 1.9 (A simple joint life status $(x : \bar{n})$ .)

Let  $(u) = (x : \bar{n})$ , with the corresponding r.v. describing the future life time being  $T(x : \bar{n})$ . The life status is a joint life status and it thus dies on the first death of either  $(x)$  or  $(\bar{n})$ . Then

$${}_t p_{x:\bar{n}} = \mathbf{P}[T(x : \bar{n}) \geq t] = \begin{cases} \mathbf{P}[T(x) \geq t] = {}_t p_x, & t < n \\ 0, & t \geq n \end{cases}$$

The r.v. is a mixed one with an atom (jump) at  $t = n$ . Also

$${}_k p_{x:\bar{n}} = \mathbf{P}[K(x : \bar{n}) \geq k] = \begin{cases} \mathbf{P}[K(x) \geq k] = {}_k p_x, & k < n \\ 0, & k \geq n \end{cases}$$

**At home.**

Show that the expectation of  $T(x : \bar{n})$  is given by

$$\mathbf{E}[T(x : \bar{n})] = \int_0^n t {}_t p_x \cdot \mu(x + t) dt + n {}_n p_x.$$

Also, check that

$$\mathbf{E}[K(x : \bar{n})] = \sum_{k=0}^{n-1} k {}_k q_x + (n-1) {}_n p_x.$$

# Joint life status.

## Example 1.10 (A joint life status $(x : y)$ .)

Let  $(u) = (x : y)$ , with the corresponding r.v. describing the future life time being  $T(x : y)$ . The life status is a joint life status and it thus dies on the first death of either  $(x)$  or  $(y)$ . Let  $T(x)$  be independent of  $T(y)$ . Then

$${}_t p_{x:y} = \mathbf{P}[T(x : y) \geq t] = {}_t p_x \cdot {}_t p_y.$$

## Ordered deaths.

Example 1.11 (A life status  $(x : \bar{n})$ .)

Let  $(u) = (x : \bar{n})$  be a life status that dies if  $(x)$  dies first (on  $[0, n)$ ), and it never dies thereafter (on  $[n, \infty)$ ) if  $(x)$  is then alive. In such a case

$${}_t p_{x:\bar{n}} = \mathbf{P}[T(x : \bar{n}) \geq t] = \begin{cases} {}_t p_x, & t < n \\ n p_x, & t \geq n \end{cases} .$$

Also

$${}_k p_{x:\bar{n}} = \mathbf{P}[K(x : \bar{n}) \geq k] = \begin{cases} {}_k p_x, & k < n \\ n p_x, & k \geq n \end{cases} .$$



**At home.**

Show that the expectation of  $T(x : \bar{n})^1$  is given by

$$\mathbf{E}[T(x : \bar{n})^1] = \infty.$$

Also, check that

$$\mathbf{E}[K(x : \bar{n})^1] = \infty.$$

### Example 1.12 (A life status $(x : \bar{n})$ .)

Let  $(u) = (x : \bar{n})$  be a life status that dies if  $(\bar{n})$  dies first. In such a case

$${}_t p_{x:\bar{n}} = \mathbf{P}[T(x : \bar{n}) \geq t] = \begin{cases} 1, & t < n \\ {}_n q_x, & t \geq n \end{cases} .$$

Also

$${}_k p_{x:\bar{n}} = \mathbf{P}[K(x : \bar{n}) \geq k] = \begin{cases} 1, & k < n \\ {}_n q_x, & k \geq n \end{cases} .$$

**At home.**

Show that the expectations with respect to the last two  $p$  functions are infinite.